

Identifying the limiting distribution by a general approach of Stein's method

In memory of Professor Xiru Chen (1934–2005)

SHAO Qi-Man* & ZHANG Zhuo-Song

Department of Statistics, The Chinese University of Hong Kong, Hong Kong, China

Email: qmshao@sta.cuhk.edu.hk, zhzhang@link.cuhk.edu.hk

Received May 12, 2016; accepted August 6, 2016; published online September 26, 2016

Abstract A general exchange pair approach is developed to identify the limiting distribution for any sequence of random variables, by calculating the conditional mean and the conditional second moments. The error of approximation is also studied. In particular, a Berry-Esseen type bound of $O(n^{-3/4})$ is obtained for the Curie-Weiss model at the critical temperature.

Keywords Stein's method, exchangeable pair, limiting distribution, Berry-Esseen bounds, Curie-Weiss model

MSC(2010) 60F05

Citation: Shao Q-M, Zhang Z-S. Identifying the limiting distribution by a general approach of Stein's method. *Sci China Math*, 2016, 59: 2379–2392, doi: 10.1007/s11425-016-0322-3

1 Introduction

Let $W := W_n$ be a sequence of random variables of interest. Since the exact distribution of W is usually unknown, it would be interesting to find out the limiting distribution. There are several approaches to solve this problem. The classical method is to calculate the characteristic function, which may not be easy to do. Another approach is to use the Stein method. Stein's method was first introduced by [18] for normal approximation. The method is striking because it can deal with not only independent random variables but also dependent random variables and it can also provide an accuracy of the approximation. Stein's idea and method have been extended to various approximation far beyond the normal approximation, for example, to Poisson approximation by Chen [5], to diffusion approximation by Barbour [1], to Gamma approximation by Luk [15], to multivariate normal approximation by Barbour [1], and Meckes and Meckes [16]. We refer to [6] for a thorough coverage of Stein methods fundamentals and recent developments in both theory and applications. We also refer to [2] for a short survey on Stein's method.

By using the exchangeable pair approach of Stein's method, Chatterjee and Shao [4] provided a concrete method to identify the limiting distribution of W under certain conditions. Let (W, W') be an exchangeable pair and $\Delta = W - W'$. Assume that there exist $\lambda > 0$, measurable functions $g(w)$, $r_1(w)$ and $r_2(w)$ such that

$$E(\Delta | W) = \lambda(g(W) + r_1(W)) \quad (1.1)$$

*Corresponding author

and

$$E(\Delta^2 | W) = 2\lambda(1 + r_2(W)). \quad (1.2)$$

Let $G(t) = \int_0^t g(s)ds$, $p(t) = c_0 \exp(-G(t))$, where $c_0 = 1/\int_{-\infty}^{\infty} e^{-G(t)}dt$. Under some regular assumptions on $g(w)$, Chatterjee and Shao [4] showed that if $E(|r_1(W)| + |r_2(W)| + |\Delta|^3/\lambda) \rightarrow 0$, then $W \xrightarrow{d} Y$, where Y has the probability density function $p(t)$.

Assumption (1.2) implies that the conditional second moment of Δ , given W , satisfies a law of large numbers. However, this may not be true in general. The main purpose of this note is to find the limiting distribution of W without the assuming condition (1.2).

The paper is organized as follows. The next section presents the main results. In Section 3, an application to Curie-Weiss model at the critical temperature is discussed with a Berry-Esseen type bound of $O(n^{-3/4})$. Section 4 provides some basic properties of Stein's equation and solution, while the proof of the main results is postponed to Section 5.

2 Main results

Let (W, W') be an exchangeable pair and $\Delta = W - W'$. Assume that there exist $\lambda > 0$ and measurable functions $g(w), v(w) \geq 0, r_1(w)$ and $r_2(w)$ such that

$$E[\Delta | W] = \lambda(g(W) + r_1(W)) \quad (2.1)$$

and

$$E[\Delta^2 | W] = 2\lambda(v(W) + r_2(W)). \quad (2.2)$$

It is well known that conditional expected value and the conditional second moment of Δ , given W , must be a measurable function of W . So Conditions (2.1) and (2.2) are always satisfied.

Let Y be a random variable with the probability density function

$$p(w) = \frac{1}{c_1 v(w)} \exp(-Q(w)), \quad w \in (a, b), \quad (2.3)$$

where $Q(w) = \int_{w_0}^w q(t)dt$, w_0 satisfies $g(w_0) = 0$, $q(t) = g(t)/v(t)$ and c_1 is the normalizing constant. Assume $v(a+)p(a+) = v(b-)p(b-) = 0$.

To present our main results, we first introduce some assumptions on the functions g and v . Assume that (B1) There exist constants $\alpha \geq 1$ and $\beta \geq 0$ such that for $w_0 \leq x \leq y < b$

$$|g(x)| \leq \alpha g(y) + \beta, \quad (2.4)$$

and $a < y \leq x \leq w_0$,

$$|g(x)| \leq -\alpha g(y) + \beta. \quad (2.5)$$

(B2) There exists a constant $c_2 \geq c_1 \max\{1, E[v(Y)]\}$ such that the equations

$$c_2 g(x) = 1, \quad c_2 g(x) = v(x) \quad (2.6)$$

have at most one solution on (w_0, b) and the equations

$$c_2 g(x) = -1, \quad c_2 g(x) = -v(x) \quad (2.7)$$

have at most one solution on (a, w_0) .

(B3) There exists an interval $[l, u] \subset (a, b)$ such that on (a, w_0) , $v(x)$ is non-decreasing or $\inf_{x \in (a, l)} v(x) \geq c_3$, on (w_0, b) , $v(x)$ is non-increasing or $\inf_{x \in (u, b)} v(x) \geq c_3$. Moreover, $\|v'\| = \sup_{x \in (a, b)} |v'(x)| \leq c_4$.

(B4) There exists a constant $c_5 \geq 1$ such that

$$\max\{1, |g'(y)|\} \min\{c_2, (\alpha + 1)(\alpha + \beta c_2)/|g(y)|\} (|y| + E|Y| + c_2) \leq c_5. \quad (2.8)$$

Remark 2.1. If g and g/v are non-decreasing on (a, b) with $(w - w_0)g(w) \geq 0$, then Conditions (B1) and (B2) are satisfied with $\alpha = 1$ and $\beta = 0$.

We are now ready to identify the limiting distribution of W .

Theorem 2.2. Let (W, W') be an exchangeable pair satisfying (2.1), (2.2) and the conditions (B1)–(B4). If

$$E|r_1(W)| + E|r_2(W)| + E|\Delta^3/(\lambda v(W))| \rightarrow 0, \tag{2.9}$$

then W converges to Y in distribution.

The next theorem gives an L_1 bound for the approximation.

Theorem 2.3. Let (W, W') be as defined in Theorem 2.2. Then for $\|h'\| < \infty$, we have

$$|E[h(W)] - E[h(Y)]| \leq C\|h'\| \left(E|r_1(W)| + E|r_2(W)| + \frac{1}{\lambda} E[|\Delta|^3/v(W)] \right), \tag{2.10}$$

where C is a finite constant depending on $w_0, c_1, \dots, c_5, \alpha$ and β .

When Δ is bounded, we can also give the Berry-Esseen bound for the approximation. Assume that

$$|\Delta| \leq \delta. \tag{2.11}$$

Also assume that

(B5) The interval (a, b) can be partitioned by three parts, I_1, I_2 and I_3 and there exists a constant δ_1 such that

$$E \sup_{|t| \leq \delta} \left| \frac{1}{v(W+t)} \right| \mathbb{1}(W \in I_1 \cup I_3) \leq \delta_1. \tag{2.12}$$

Moreover, v is absolutely continuous and there exist constants c_6 and c_7 such that

$$\sup_{|t| \leq \delta} \sup_{x \in I_2} \left| \frac{1}{v(x+t)} \right| \leq c_6, \quad \sup_{|t| \leq \delta} \sup_{x \in I_2} |v'(x+t)| \leq c_7. \tag{2.13}$$

We have the following Berry-Essen type inequality:

Theorem 2.4. Assume that (2.1), (2.2), (2.11) and Conditions (B1), (B2), (B4) and (B5) are satisfied. Then

$$\begin{aligned} & \sup_{z \in (a,b)} |P(W \leq z) - P(Y \leq z)| \\ & \leq 2c_2 E|r_1(W)| + E \left| \frac{r_2(W)}{v(W)} \right| + \frac{2(\alpha + \beta c_2 + 1)\delta^2 \delta_1}{\lambda} + (\|p\| + 2c_5/c_2)\delta \\ & \quad + \frac{\delta^3}{\lambda} ((\alpha + \beta c_2)c_6^2 c_7 + c_5 c_6/c_2 + c_5 c_6 E|g(W)| + c_6^2(\alpha + 1 + \beta c_2)E|g(W)|). \end{aligned} \tag{2.14}$$

3 Application to Curie-Weiss model

Consider the Curie-Weiss model for n -spins $\sigma = (\sigma_1, \dots, \sigma_n) \in \{-1, 1\}^n$ at temperature T . The joint density function of σ is given by

$$\frac{1}{Z_T} \exp \left(\frac{\sum_{1 \leq i < j \leq n} \sigma_i \sigma_j}{Tn} \right), \tag{3.1}$$

where Z_T is the normalizing constant. For the critical temperature $T = 1$, let

$$W = W(\sigma) = n^{-3/4} \sum_{i=1}^n \sigma_i. \tag{3.2}$$

This is a simple statistical mechanical model of ferromagnetic interaction, also called the Ising model on the complete graph. For a detailed mathematical treatment of this model, we refer to the book by [8].

It was proved by [9–11] that as $n \rightarrow \infty$, the law of W converges to the distribution with density proportional to $e^{-x^4/12}$ and Chatterjee and Shao [4] obtained a Berry-Esseen bound of $O(n^{-1/2})$. For various interesting extensions and refinements of their results, one can refer to [12, 17].

In this section, we shall prove that the Berry-Esseen bound can be improved to $O(n^{-3/4})$ when the “limiting distribution” is allowed to depend on n , which in turn also shows that the result obtained by Chatterjee and Shao [4] is optimal.

We first construct W' so that (W, W') is an exchangeable pair. Let I be a uniformly distributed random index over $\{1, \dots, n\}$. For each i , given $\sigma_j, j \neq i, 1 \leq j \leq n$, let σ'_i be independent of σ_i and have the same conditional distribution as σ_i . Set $W' = W(\sigma_1, \dots, \sigma_{I-1}, \sigma'_I, \sigma_{I+1}, \dots, \sigma_n)$. Then (W, W') is an exchangeable pair. The following lemma verifies various conditions in Theorems 2.3 and 2.4.

Lemma 3.1. *Let $\Delta = W - W'$. We have,*

$$E[\Delta | W] = \frac{1}{3}n^{-3/2}W^3 - n^{-2}W + O(n^{-5/2})(1 + W^5), \quad (3.3)$$

$$E[\Delta^2 | W] = 2n^{-3/2} \max\{1 - n^{-1/2}W^2, n^{-1}\} + O(n^{-5/2})(1 + W^4), \quad (3.4)$$

$$|W - W'| \leq 2n^{-3/4} \quad (3.5)$$

and

$$E|W|^3 \leq 15. \quad (3.6)$$

From this lemma, we can choose $\lambda = n^{-3/2}$,

$$g(w) = \frac{w^3}{3} - n^{-1/2}w, \quad v(w) = \max\{1 - n^{-1/2}w^2, n^{-1}\},$$

$$|r_1(w)| \leq An^{-1}|w|^3, \quad |r_2(w)| \leq An^{-1}(1 + w^4),$$

where A is an absolute constant.

Let

$$q(w) = \frac{w^3/3 - n^{-1/2}w}{\max\{1 - n^{-1/2}w^2, n^{-1}\}},$$

$$Q(y) = \int_0^y q(w)dw, \quad (3.7)$$

and Y be a random variable with the probability density function

$$p(y) = \frac{1}{c_1 v(y)} e^{-Q(y)}, \quad y \in (-\infty, \infty), \quad (3.8)$$

where c_1 is the normalizing constant.

Theorem 3.2. *We have for any absolutely continuous function h with $\|h'\| < \infty$,*

$$|E[h(W) - h(Y)]| \leq C\|h'\|n^{-3/4} \quad (3.9)$$

and

$$\sup_z |P(W \leq z) - P(Y \leq z)| \leq Cn^{-3/4}, \quad (3.10)$$

where C is an absolute constant.

Remark 3.3. From (3.10) and (3.7), we see that $P(Y \leq z)$ involves a term of order $n^{-1/2}$. This indicates that the error bound of order $O(n^{-1/2})$ given by Chatterjee and Shao [4] is optimal.

Proof. By Theorems 2.3 and 2.4, it suffices to show that Conditions (B1)–(B5) are satisfied. It is easy to see that for all $0 \leq x \leq y \leq n^{1/4}$, $|g(x)| \leq g(y) + 1$ and for $0 \geq x \geq y \geq -n^{1/4}$, $|g(x)| \leq -g(y) + 1$. It is also not difficult to verify that Conditions (B2)–(B4) are satisfied.

As for (B5), we choose $I_1 = (-\infty, -n^{1/4}/\sqrt{2})$, $I_2 = [-n^{1/4}/\sqrt{2}, n^{1/4}/\sqrt{2}]$ and $I_3 = (n^{1/4}/\sqrt{2}, \infty)$. Recall that $v(w) = \max\{1 - n^{-1/2}w^2, n^{-1}\}$, then for $w \in I_1 \cup I_3$, $|t| \leq \delta$, $v(w + t) \geq n^{-1}$.

By [3, Proposition 4], for $t \geq 0$, $P(|W| \geq t) \leq 2e^{-ct^4}$, where $c > 0$ is an absolute constant, we have by integration by parts,

$$\begin{aligned} & \mathbb{E} \left[\sup_{|t| \leq \delta} \left| \frac{1}{v(W+t)} \right| \cdot \mathbb{1}(n^{-1/2}W^2 > 1/2) \right] \\ & \leq nP(W^2 > n^{1/2}/2) + n \int_{n^{1/2}/2}^{\infty} P(W^2 \geq y) dy \\ & \leq Cn^{-3/4}. \end{aligned}$$

Similarly, we can prove that

$$\mathbb{E} \left| \frac{1}{v(W)} \right| \leq C, \quad \mathbb{E}|r_1(W)| \leq Cn^{-1}, \quad \mathbb{E}|r_2(W)| \leq Cn^{-1} \quad \text{and} \quad \mathbb{E} \left| \frac{r_2(W)}{v(W)} \right| \leq Cn^{-1}$$

for some absolute constant C . One can also check that $\|p\| \leq C$ for some constant C . This completes the proof. \square

We now turn to the proof of Lemma 3.1.

Proof of Lemma 3.1. Let

$$M = n^{-1} \sum_{i=1}^n \sigma_i, \quad M_i = n^{-1} \sum_{j \neq i} \sigma_j.$$

Thus $|M| \leq 1$ and $|W| \leq n^{-1/4}$.

Let \mathcal{F} be the sigma field generated by σ . By [4], we have

$$\begin{aligned} \mathbb{E}[\Delta | \mathcal{F}] &= n^{-3/4}M - n^{-7/4} \sum_{i=1}^n \tanh M_i \\ &= n^{-3/4}M - n^{-7/4} \sum_{i=1}^n \left(M_i - \frac{M_i^3}{3} + O(1)M_i^5 \right) \\ &= n^{-3/2} \left(\frac{W^3}{3} - n^{-1/2}W + O(n^{-1})(1 + |W|^5) \right). \end{aligned} \tag{3.11}$$

This proves (3.3).

By [4] again, we have

$$\begin{aligned} \mathbb{E}[\Delta^2 | \mathcal{F}] &= 2n^{-3/2} - 2n^{-5/2} \sum_{i=1}^n \sigma_i \tanh M_i \\ &= 2n^{-3/2} - 2n^{-5/2} \sum_{i=1}^n \sigma_i \left(M_i - \frac{M_i^3}{3} + O(1)M_i^5 \right) \\ &= 2n^{-3/2}(1 - n^{-1}W^2 + O(n^{-1}(1 + W^4))) \\ &= 2n^{-3/2} \max\{1 - n^{-1}W^2, n^{-1}\} + O(n^{-5/2}(1 + W^4)). \end{aligned} \tag{3.12}$$

The inequalities (3.5) and (3.6) follows directly from [4]. This completes the proof of Lemma 3.1. \square

4 Properties of the Stein solution

In this section, we will recall or prove some basic properties of the Stein solution, mainly following the arguments in [4]. We start with the Stein equation.

Let Y be a random variable with the probability density function (2.3). For any absolutely continuous function f , one can show that

$$\mathbb{E}[v(Y)f'(Y)] = \mathbb{E}[g(Y)f(Y)]. \quad (4.1)$$

For a given measurable function h , let f_h be the solution to the following Stein equation:

$$v(w)f'(w) - g(w)f(w) = h(w) - \mathbb{E}[h(Y)], \quad w \in (a, b). \quad (4.2)$$

It is known that the solution f_h can be expressed as for $w \in (a, b)$,

$$\begin{aligned} f_h(w) &= \frac{1}{v(w)p(w)} \int_a^w (h(t) - \mathbb{E}[h(Y)])p(t)dt \\ &= -\frac{1}{v(w)p(w)} \int_w^b (h(t) - \mathbb{E}[h(Y)])p(t)dt. \end{aligned} \quad (4.3)$$

Equation (4.2) has also been discussed in [7, 13, 14], where an error bound of certain distances (but not including L_1 distance) of beta distribution approximation were obtained. Theorem 2.3 provides an L_1 bound for general non-normal approximation.

Lemma 4.1. *Assume that (B4) is satisfied and $\delta > 0$ with $(c_5/c_2)\delta \leq 1/2$. Then we have*

$$\sup_{|t| \leq \delta} |g(w+t) - g(w)| \leq 2(c_5/c_2^2)\delta + 2(c_5/c_2)\delta|g(w)|. \quad (4.4)$$

Proof. From (2.8) it follows that

$$|g'(w)| \leq (c_5/c_2)(1/c_2 + |g(w)|). \quad (4.5)$$

Thus by the mean value theorem,

$$\begin{aligned} \sup_{|t| \leq \delta} |g(w+t) - g(w)| &\leq \delta \sup_{|t| \leq \delta} |g'(w+t)| \\ &\leq \frac{c_5}{c_2} \delta (1/c_2 + |g(w+t)|) \\ &\leq (c_5/c_2^2)\delta + (c_5/c_2)\delta|g(w)| + (c_5/c_2)\delta \sup_{|t| \leq \delta} |g(w+t) - g(w)| \\ &\leq (c_5/c_2^2)\delta + (c_5/c_2)\delta|g(w)| + \frac{1}{2} \sup_{|t| \leq \delta} |g(w+t) - g(w)|. \end{aligned}$$

This yields (4.4). □

Lemma 4.2. *Let $\|h\| < \infty$ and f_h be the solution to the equation (4.2). Under Conditions (B1)–(B4), we have*

$$\|f_h\| \leq 2c_2\|h\|, \quad \|f_h g\| \leq 2(\alpha + \beta c_2)\|h\|. \quad (4.6)$$

Proof. Recall that the solution to (4.2) is given by

$$\begin{aligned} f_h(w) &= \frac{1}{v(w)p(w)} \int_a^w (h(t) - \mathbb{E}[h(Y)])p(t)dt \\ &= -\frac{1}{v(w)p(w)} \int_w^b (h(t) - \mathbb{E}[h(Y)])p(t)dt. \end{aligned}$$

For $w \leq w_0$, define

$$H(w) = \int_a^w p(t)dt - c_2 v(w)p(w)$$

and thus

$$H'(w) = p(w)(1 + c_2 g(w)).$$

Since $g(w_0) = 0$ and there exists at most one point w_1 such that $c_2 g(w_1) + 1 = 0$, $H(w)$ achieves its maximum at w_0 or a . Note that $H(a) \leq 0$ by definition and

$$H(w_0) \leq 1 - c_2/c_1 \leq 0.$$

This gives $H(w) \leq 0$ for $w \in (a, w_0]$ and thus $|f_h(w)| \leq c_2 \|h\|$. Similarly, we can prove $f_h(w) \leq c_2 \|h\|$ on (w_0, b) .

Next, we will give the bound of $\|f_h g\|$. For $w \in (w_0, b)$, by (2.4),

$$\begin{aligned} |f_h(w)g(w)| &= |g(w)|e^{Q(w)} \int_w^b \frac{1}{v(t)} e^{-Q(t)} dt \\ &\leq \alpha e^{Q(w)} \int_w^b \frac{g(t)}{v(t)} e^{-Q(t)} dt + \beta e^{Q(w)} \int_w^b p(t) dt \\ &\leq \alpha + \beta c_2. \end{aligned}$$

Similarly, for $w \in (a, w_0)$,

$$|f_h(w)g(w)| \leq \alpha + \beta c_2.$$

This completes the proof. □

Lemma 4.3. *Let f_h be the solution given by (4.3). Assume that Conditions (B1)–(B4) are satisfied and that h is absolutely continuous with $\|h'\| < \infty$. Then*

$$\|f_h\| \leq c_5 \|h'\|, \quad \|f'_h\| \leq C \|h'\|, \tag{4.7}$$

$$\|f_h g'\| \leq c_5 \|h'\|, \quad \|f'_h g\| \leq C \|h'\|, \tag{4.8}$$

where C is a constant depending on $\alpha, \beta, c_2, c_3, c_4$ and c_5 .

Proof. Since h is absolutely continuous,

$$h(y) - E[h(Y)] = \int_a^y h'(t)F(t)dt - \int_y^b h'(t)(1 - F(t))dt.$$

By (4.3), we have

$$\begin{aligned} f_h(y)v(y)p(y) &= \int_a^y (h(t) - E[h(Y)])p(t)dt \\ &= -(1 - F(y)) \int_a^y h'(t)F(t)dt - F(y) \int_y^b h'(t)(1 - F(t))dt. \end{aligned} \tag{4.9}$$

Hence,

$$|f(y)v(y)p(y)| \leq \|h'\| \left((1 - F(y)) \int_a^y F(t)dt + F(y) \int_y^b (1 - F(t))dt \right). \tag{4.10}$$

From (2.6), we know for $y \in (w_0, b)$,

$$1 - F(y) \leq c_2 v(y)p(y) \min\{1, 1/|g(y)|\}.$$

By Condition (B2), we can similarly prove

$$\int_y^b v(t)p(t)dt \leq c_2 v(y)p(y), \quad \text{for } y \in (a, b).$$

Note that for $y \in (w_0, b)$,

$$\int_a^y F(s)ds = yF(y) - \int_a^y tp(t)dt. \quad (4.11)$$

Thus,

$$\begin{aligned} & (1 - F(y)) \int_a^y F(t)dt + F(y) \int_y^b (1 - F(t))dt \\ & \leq \min\{c_2, (\alpha + \beta c_2)/|g(y)|\}(|y| + E|Y|)v(y)p(y) + \int_y^b \min\{c_2, (\alpha + \beta c_2)/|g(t)|\}v(t)p(t)dt \\ & \leq \min\{c_2, (\alpha + \beta c_2)/|g(y)|\}(|y| + E|Y|)v(y)p(y) + \min\{c_2, \alpha(\alpha + \beta c_2)/(|g(y)| - \beta)\} \int_y^b v(t)p(t)dt, \end{aligned}$$

where we used (2.4) in the last line. When $|g(y)| - \beta \leq \alpha^2/c_2 + \alpha\beta$, we have $\min\{c_2, \alpha(\alpha + \beta c_2)/(|g(y)| - \beta)\} = c_2$. Otherwise, $|g(y)| \geq (\alpha + 1)\beta$ and thus

$$\alpha(\alpha + \beta c_2)/(|g(y)| - \beta) \leq (\alpha + 1)(\alpha + \beta c_2)/|g(y)|.$$

Hence, we can rewrite these terms above as

$$\begin{aligned} & (1 - F(y)) \int_a^y F(t)dt + F(y) \int_y^b (1 - F(t))dt \\ & \leq \min\{c_2, (\alpha + \beta c_2)/|g(y)|\}(|y| + E|Y|)v(y)p(y) \\ & \quad + \min\{c_2, (\alpha + 1)(\alpha + \beta c_2)/|g(y)|\}c_2 v(y)p(y). \end{aligned} \quad (4.12)$$

Hence, by (2.8), for $y \in (w_0, b)$,

$$\begin{aligned} & \max\{1, |g'(y)|\} \left((1 - F(y)) \int_a^y F(t)dt + F(y) \int_y^b (1 - F(t))dt \right) \\ & \leq v(y)p(y) \max\{1, |g'(y)|\} \min\{c_2, (\alpha + 1)(\alpha + \beta c_2)/|g(y)|\}(|y| + E|Y| + c_2) \\ & \leq c_5 v(y)p(y). \end{aligned} \quad (4.13)$$

Similarly, for $y \in (a, w_0)$,

$$\max\{1, |g'(y)|\} \left((1 - F(y)) \int_a^y F(t)dt + F(y) \int_y^b (1 - F(t))dt \right) \leq c_5 v(y)p(y). \quad (4.14)$$

By (4.10), (4.13) and (4.14), we have

$$\|f_h\| \leq c_5 \|h'\|, \quad \|f_h g'\| \leq c_5 \|h'\|. \quad (4.15)$$

Next, we prove the second inequality of (4.7). By [7], we have

$$|f'_h(y)| \leq \|h'\| \left(\frac{\int_a^y F(s)ds |G(y)| + \int_y^b (1 - F(s))ds |H(y)|}{v^2(y)p(y)} \right), \quad (4.16)$$

where

$$G(y) = v(y)p(y) - g(y)(1 - F(y)), \quad H(y) = v(y)p(y) + g(y)F(y).$$

When $y \in (w_0, b)$ and if v is non-increasing,

$$\begin{aligned} & \left| \int_y^b (1 - F(s)) ds \right| + \frac{|g(y)F(y)|}{v(y)p(y)} \\ & \leq c_2 \int_y^b v(s)p(s) ds \left| 1 + \frac{|g(y)|}{v(y)p(y)} \right| \\ & \leq c_2(1 + \alpha + \beta c_2)v(y) \end{aligned}$$

and thus

$$\frac{\int_y^b (1 - F(s)) ds |H(y)|}{v^2(y)p(y)} \leq c_2(1 + \alpha + \beta c_2). \tag{4.17}$$

By (4.11),

$$\frac{\int_a^y F(s) ds |G(y)|}{v^2(y)p(y)} \leq \frac{|y|G(y)}{v^2(y)p(y)} + \frac{E|Y|G(y)}{v^2(y)p(y)}.$$

Recall that

$$v(y)p(y) = - \int_a^y g(t)p(t) dt = \int_y^b g(t)p(t) dt,$$

then

$$\begin{aligned} \frac{(|y| + E|Y|)|G(y)|}{v^2(y)p(y)} &= \frac{(|y| + E|Y|) \left| \int_y^b (g(t) - g(y))p(t) dt \right|}{v^2(y)p(y)} \\ &= \frac{(|y| + E|Y|) \left| \int_y^b g'(t)(1 - F(t)) dt \right|}{v^2(y)p(y)} \\ &\leq \frac{c_2(|y| + E|Y|) \int_y^b |g'(t)|v(t)p(t) dt}{v^2(y)p(y)} \\ &\leq \frac{c_2 \int_y^b |g'(t)|(|t| + E|Y|)p(t) dt}{v(y)p(y)} \\ &\leq \frac{c_5 \int_y^b (|g(t)| + 1/c_2)p(t) dt}{v(y)p(y)} \\ &\leq \frac{c_5 \int_y^b (\alpha g(t) + 1/c_2 + \beta)p(t) dt}{v(y)p(y)} \\ &\leq C, \end{aligned}$$

where C is a constant depending on c_2, c_5, α and β . Therefore,

$$\sup_{y \in (w_0, b)} \frac{\int_a^y F(s) ds |G(y)| + \int_y^b (1 - F(s)) ds |H(y)|}{v^2(y)p(y)} \leq C.$$

Similarly, we can also prove that if $1/v(y) \leq 1/c_3$ we have

$$\sup_{y \in (w_0, b)} \frac{\int_a^y F(s) ds |G(y)| + \int_y^b (1 - F(s)) ds |H(y)|}{v^2(y)p(y)} \leq C.$$

For $y \in (a, w_0)$, if $v(y)$ is non-decreasing or $1/v(y) \leq 1/c_3$, we also have

$$\sup_{y \in (a, w_0)} \frac{\int_a^y F(s) ds |G(y)| + \int_y^b (1 - F(s)) ds |H(y)|}{v^2(y)p(y)} \leq C.$$

This proves $\|f'_h\| \leq C\|h'\|$.

Finally, we will give the bound of $f'_h g$. From (4.2), we have

$$f''_h v - f'_h g = h' + g' f_h - v' f'_h.$$

Thus

$$(f'_h v p)' = f''_h v p - f'_h g p = h' p + g' f_h p - v' f'_h p$$

and

$$\begin{aligned} |g(x) f'_h(x) v(x) p(x)| &= \left| g(x) \int_a^x (h' p)(t) + (g' f_h)(t) - (v' f'_h p)(t) dt \right| \\ &= \left| g(x) \int_x^b (h' p)(t) + (g' f_h p)(t) - (v' f'_h p)(t) dt \right| \\ &\leq C \|h'\| |g(x)| \min\{F(x), 1 - F(x)\} \\ &\leq C \|h'\| v(x) p(x). \end{aligned}$$

This proves $\|f'_h g\| \leq C\|h'\|$ where C is a constant. \square

Lemma 4.4. *Let f_h be the solution given by (4.3) and satisfy the conditions in Lemma 4.3. We have*

$$|f'_h(x+t) - f'_h(x)| \leq C \|h'\| |t|/v(x). \quad (4.18)$$

Proof. Observe that

$$\begin{aligned} &|f'_h(x+t) - f'_h(x)| \\ &= \left| \frac{f_h(x+t)g(x+t)}{v(x+t)} - \frac{f_h(x)g(x)}{v(x)} + \frac{h(x+t) - \mathbb{E}[h(Y)]}{v(x+t)} - \frac{h(x) - \mathbb{E}[h(Y)]}{v(x)} \right| \\ &\leq |f_h(x+t)g(x+t) + h(x+t) - \mathbb{E}h(Y)| \times \left| \frac{1}{v(x+t)} - \frac{1}{v(x)} \right| \\ &\quad + \frac{1}{v(x)} |f_h(x+t)g(x+t) - f_h(x)g(x)| + \frac{1}{v(x)} |h(x+t) - h(x)| \\ &=: L_1 + L_2 + L_3. \end{aligned}$$

We next give the bounds of L_1 , L_2 and L_3 . For L_1 ,

$$\left| \frac{1}{v(x+t)} - \frac{1}{v(x)} \right| \leq \frac{|v(x+t) - v(x)|}{v(x)v(x+t)} \leq \frac{c_4 |t|}{v(x+t)v(x)}$$

and by (4.6), we have

$$L_1 \leq \frac{c_4 |f'_h(x+t)| |t|}{v(x)} \leq C \|h'\| |t|/v(x). \quad (4.19)$$

For L_2 , it is easy to see

$$L_2 \leq \frac{(\|f_h g'\| + \|f'_h g\|) |t|}{v(x)} \leq C \|h'\| |t|/v(x). \quad (4.20)$$

Finally, for L_3 , we have $L_3 \leq \|h'\| |t|/v(x)$. This proves the lemma. \square

5 Proof of main results

Theorem 2.2 is a direct consequence of Theorem 2.3.

Proof of Theorem 2.3. Recall that $\Delta = W - W'$ and observe that

$$\begin{aligned} 0 &= \mathbb{E}[(W - W')(f(W) + f(W'))] \\ &= 2\lambda\mathbb{E}[f(W)g(W)] + 2\lambda\mathbb{E}[f(W)r_1(W)] - \mathbb{E}[\Delta(f(W) - f(W'))] \\ &= 2\mathbb{E}[f(W)\mathbb{E}[\Delta | W]] - \mathbb{E}\left[\Delta \int_{-\Delta}^0 f'(W + t)dt\right] \\ &= 2\lambda\mathbb{E}[f(W)g(W)] + 2\lambda\mathbb{E}[f(W)r_1(W)] - 2\lambda\mathbb{E}\left[\int_{-\infty}^{\infty} f'(W + t)\hat{K}(t)dt\right], \end{aligned}$$

where $\hat{K}(t) = \mathbb{E}[\frac{\Delta}{2\lambda}(\mathbf{1}(-\Delta \leq t \leq 0) - \mathbf{1}(0 \leq t \leq -\Delta)) | W]$. Therefore,

$$\mathbb{E}[f(W)g(W)] = \mathbb{E}\int_{-\infty}^{\infty} f'(W + t)\hat{K}(t)dt - \mathbb{E}[f(W)r_1(W)]. \tag{5.1}$$

By (2.2), we have

$$\frac{1}{2\lambda}\mathbb{E}[f'(W)\Delta^2] = \mathbb{E}[f'(W)v(W)] + \mathbb{E}[f'(W)r_2(W)]. \tag{5.2}$$

Now for f_h given in (4.3), by (5.1), (5.2), Lemmas 4.3 and 4.4, we have

$$\begin{aligned} |\mathbb{E}[h(W)] - \mathbb{E}[h(Y)]| &= |\mathbb{E}[v(W)f'_h(W) - g(W)f_h(W)]| \\ &= \left| \mathbb{E}[f_h(W)r_1(W)] - \mathbb{E}[f'_h(W)r_2(W)] - \mathbb{E}\int_{-\infty}^{\infty} f'(W + t) - f'(W)dt \right| \\ &\leq C\|h'\| \left(\mathbb{E}|r_1(W)| + \mathbb{E}|r_2(W)| + \mathbb{E}\left[\frac{1}{v(W)}\int_{-\infty}^{\infty} |t|\hat{K}(t)dt\right] \right) \\ &\leq C\|h'\|(\mathbb{E}|r_1(W)| + \mathbb{E}|r_2(W)| + \mathbb{E}|\Delta^3/\lambda v(W)|). \end{aligned} \tag{5.3}$$

This completes the proof. □

Proof of Theorem 2.4. The inequality (2.14) is trivial if $2c_5\delta/c_2 > 1$. We assume that $2c_5\delta/c_2 \leq 1$ below. Now let $h(W) = \mathbf{1}(W \leq z)$. Assume that $|\Delta| \leq \delta$. By (5.1),

$$\begin{aligned} \mathbb{E}[f(W)g(W) + f(W)r_1(W)] &= \mathbb{E}\left[\int_{|t|\leq\delta} f'(W + t)\hat{K}(t)dt\right] \\ &= \mathbb{E}\left[\int_{|t|\leq\delta} f(W + t)q(W + t)\hat{K}(t)dt\right] \\ &\quad + \mathbb{E}\left[\int_{|t|\leq\delta} \frac{(h(W + t) - \mathbb{E}[h(Y)])\hat{K}(t)}{v(W + t)}dt\right]. \end{aligned} \tag{5.4}$$

Observe that

$$\begin{aligned} \mathbb{E}\left[\int_{|t|\leq\delta} \frac{(h(W + t) - \mathbb{E}[h(Y)])\hat{K}(t)}{v(W + t)}dt\right] &\geq \mathbb{E}\left[\left(\frac{\mathbf{1}(W \leq z - \delta) - \mathbb{P}(Y \leq z)}{v(W)}\right)\int_{|t|\leq\delta} \hat{K}(t)dt\right] \\ &\quad - \left| \mathbb{E}\left[\int_{|t|\leq\delta} \left(\frac{1}{v(W)} - \frac{1}{v(W + t)}\right)\hat{K}(t)dt\right] \right| \\ &\geq \mathbb{P}(W \leq z - \delta) - \mathbb{P}(Y \leq z) - \mathbb{E}\left|\frac{r_2(W)}{v(W)}\right| \\ &\quad - \left| \mathbb{E}\left[\int_{|t|\leq\delta} \left(\frac{1}{v(W)} - \frac{1}{v(W + t)}\right)\hat{K}(t)dt\right] \right|. \end{aligned} \tag{5.5}$$

Then we have by (5.4),

$$\mathbb{P}(W \leq z - \delta) - \mathbb{P}(Y \leq z) \leq \mathbb{E}\left[\int_{|t|\leq\delta} (f(W)q(W) - f(W + t)q(W + t))\hat{K}(t)dt\right]$$

$$\begin{aligned}
& + \left| \mathbb{E} \left[\int_{|t| \leq \delta} \left(\frac{1}{v(W)} - \frac{1}{v(W+t)} \right) \hat{K}(t) dt \right] \right| \\
& + 2c_2 \mathbb{E} |r_1(W)| + \mathbb{E} \left| \frac{r_2(W)}{v(W)} \right|,
\end{aligned} \tag{5.6}$$

where we used (4.6) in the last inequality.

We only need to obtain the bounds of the following two terms:

$$\mathbb{E} \left[\int_{|t| \leq \delta} \left(\frac{1}{v(W)} - \frac{1}{v(W+t)} \right) \hat{K}(t) dt \right] \tag{5.7}$$

and

$$\mathbb{E} \left[\int_{|t| \leq \delta} (f(W)q(W) - f(W+t)q(W+t)) \hat{K}(t) dt \right]. \tag{5.8}$$

For the term (5.7), we have

$$\begin{aligned}
& \mathbb{E} \left[\int_{|t| \leq \delta} \left(\frac{1}{v(W)} - \frac{1}{v(W+t)} \right) \hat{K}(t) dt \right] \\
& \leq \frac{\delta^2}{\lambda} \mathbb{E} \left[\sup_{|t| \leq \delta} \left| \frac{1}{v(W+t)} \right| \mathbf{1}(W \in I_1 \cup I_3) \right] \\
& \quad + \mathbb{E} \left[\int_{|t| \leq \delta} \left(\frac{1}{v(W)} - \frac{1}{v(W+t)} \right) \hat{K}(t) dt \mathbf{1}(W \in I_2) \right] \\
& \leq \frac{\delta^2 \delta_1}{\lambda} + \frac{c_6^2 c_7 \delta^3}{4\lambda}.
\end{aligned}$$

As to (5.8), we first find the bound of $\sup_{|t| \leq \delta} \left| \frac{f(W+t)g(W+t)}{v(W+t)} - \frac{f(W)g(W)}{v(W)} \right| \cdot \mathbf{1}(x \in I_2)$. Note that

$$\begin{aligned}
& \left| \frac{f(W+t)g(W+t)}{v(W+t)} - \frac{f(W)g(W)}{v(W)} \right| \\
& \leq \left| \frac{f(W+t)g(W+t)}{v(W+t)} - \frac{f(W+t)g(W+t)}{v(W)} \right| \\
& \quad + \left| \frac{f(W+t)g(W+t)}{v(W)} - \frac{f(W+t)g(W)}{v(W)} \right| \\
& \quad + \left| \frac{f(W+t)g(W)}{v(W)} - \frac{f(W)g(W)}{v(W)} \right| \\
& =: J_1 + J_2 + J_3.
\end{aligned}$$

Recalling that

$$\sup_{|t| \leq \delta} \sup_{x \in I_2} \left| \frac{1}{v(x+t)} \right| \leq c_6, \quad \sup_{|t| \leq \delta} \sup_{x \in I_2} |v'(x+t)| \leq c_7,$$

we have

$$J_1 \leq \frac{\|fg\|c_4|t|}{v(W)v(W+t)} \leq c_6^2 c_7 \|fg\| |t|.$$

For J_2 , we have for $x \in I_2$ and $c_5\delta/c_2 \leq 1/2$,

$$\sup_{|t| \leq \delta} |g(x+t) - g(x)| \leq 2(c_5/c_2^2)\delta + 2(c_5/c_2)\delta|g(w)|,$$

then $J_2 \leq 2(c_5c_6/c_2)\|f\|\delta(1/c_2 + |g(W)|)$. For J_3 ,

$$J_3 \leq c_6\delta \sup_{|t| \leq \delta} |f'(W+t)g(W)|$$

$$\leq c_6^2(\|fg\| + 1)\delta|g(W)|.$$

Therefore, recalling that $\|f\| \leq 2c_2$, $\|fg\| \leq 2(\alpha + \beta c_2)$ we have

$$\begin{aligned} & \mathbb{E} \left[\int_{|t| \leq \delta} (f(W)q(W) - f(W+t)q(W+t))\hat{K}(t)dt \cdot \mathbf{1}(W \in I_2) \right] \\ & \leq \frac{\delta^3}{\lambda} ((\alpha + \beta c_2)c_6^2 c_7 + c_5 c_6 / c_2 + c_5 c_6 \mathbb{E}|g(W)| + c_6^2(\alpha + 1 + \beta c_2)\mathbb{E}|g(W)|). \end{aligned}$$

For (5.8), similarly we have

$$\begin{aligned} & \mathbb{E} \left[\int_{|t| \leq \delta} (f(W)q(W) - f(W+t)q(W+t))\hat{K}(t)dt \right] \\ & \leq \mathbb{E} \left[\int_{|t| \leq \delta} |f(W)q(W) - f(W+t)q(W+t)|\hat{K}(t)dt \mathbf{1}(W \in I_1 \cup I_3) \right] \\ & \quad + \mathbb{E} \left[\int_{|t| \leq \delta} |f(W)q(W) - f(W+t)q(W+t)|\hat{K}(t)dt \mathbf{1}(W \in I_2) \right] \\ & \leq \frac{\|fg\|\delta^2}{\lambda} \mathbb{E} \left[\sup_{|t| \leq \delta} \left| \frac{1}{v(W+t)} \right| \right] + \mathbb{E} \left[\int_{|t| \leq \delta} |f(W)q(W) - f(W+t)q(W+t)|\hat{K}(t)dt \mathbf{1}(W \in I_2) \right] \\ & \leq \frac{2(\alpha + \beta c_2)\delta^2 \delta_1}{\lambda} + \frac{\delta^3}{\lambda} ((\alpha + \beta c_2)c_6^2 c_7 + c_5 c_6 / c_2 + c_5 c_6 \mathbb{E}|g(W)| + c_6^2(\alpha + 1 + \beta c_2)\mathbb{E}|g(W)|). \end{aligned}$$

Therefore, combining those inequalities, we have

$$\begin{aligned} & \mathbb{P}(W \leq z - \delta) - \mathbb{P}(Y \leq z) \\ & \leq 2c_2 \mathbb{E}|r_1(W)| + \mathbb{E} \left| \frac{r_2(W)}{v(W)} \right| + \frac{2(\alpha + \beta c_2 + 1)\delta^2 \delta_1}{\lambda} \\ & \quad + \frac{\delta^3}{\lambda} ((\alpha + \beta c_2)c_6^2 c_7 + c_5 c_6 / c_2 + c_5 c_6 \mathbb{E}|g(W)| + c_6^2(\alpha + 1 + \beta c_2)\mathbb{E}|g(W)|). \end{aligned} \tag{5.9}$$

Moreover, $\mathbb{P}(z - \delta \leq Y \leq z) \leq \|p\|\delta$. Hence,

$$\begin{aligned} & \mathbb{P}(W \leq z - \delta) - \mathbb{P}(Y \leq z - \delta) \\ & \leq 2c_2 \mathbb{E}|r_1(W)| + \mathbb{E} \left| \frac{r_2(W)}{v(W)} \right| + \frac{2(\alpha + \beta c_2 + 1)\delta^2 \delta_1}{\lambda} + \|p\|\delta \\ & \quad + \frac{\delta^3}{\lambda} ((\alpha + \beta c_2)c_6^2 c_7 + c_5 c_6 / c_2 + c_5 c_6 \mathbb{E}|g(W)| + c_6^2(\alpha + 1 + \beta c_2)\mathbb{E}|g(W)|). \end{aligned}$$

Similarly,

$$\begin{aligned} & \mathbb{P}(Y \leq z + \delta) - \mathbb{P}(W \leq z + \delta) \\ & \leq 2c_2 \mathbb{E}|r_1(W)| + \mathbb{E} \left| \frac{r_2(W)}{v(W)} \right| + \frac{2(\alpha + \beta c_2 + 1)\delta^2 \delta_1}{\lambda} + \|p\|\delta \\ & \quad + \frac{\delta^3}{\lambda} ((\alpha + \beta c_2)c_6^2 c_7 + c_5 c_6 / c_2 + c_5 c_6 \mathbb{E}|g(W)| + c_6^2(\alpha + 1 + \beta c_2)\mathbb{E}|g(W)|). \end{aligned}$$

This completes the proof. □

Acknowledgements This work was supported by Hong Kong Research Grants Council General Research Fund (Grant Nos. 403513 and 14302515).

References

1 Barbour A D. Stein’s method for diffusion approximations. *Probab Theory Related Fields*, 1990, 84: 297–322

- 2 Chatterjee S. A short survey of Stein's method. ArXiv:1404.1392, 2014
- 3 Chatterjee S, Dey P S. Applications of Stein's method for concentration inequalities. *Ann Probab*, 2010, 38: 2443–2485
- 4 Chatterjee S, Shao Q-M. Nonnormal approximation by Stein's method of exchangeable pairs with application to the Curie-Weiss model. *Ann Appl Probab*, 2011, 21: 464–483
- 5 Chen L H Y. Poisson approximation for dependent trials. *Ann Probab*, 1975, 3: 534–545
- 6 Chen L H Y, Goldstein L, Shao Q-M. Normal approximation by Stein's method. *Probab Appl*, 2011, 23: 15–29
- 7 Döbler C. Stein's method of exchangeable pairs for the Beta distribution and generalizations. ArXiv:1411.4477, 2014
- 8 Ellis R S. *Entropy, Large Deviations, and Statistical Mechanics*. New York: Springer, 2012
- 9 Ellis R S, Newman C M. Limit theorems for sums of dependent random variables occurring in statistical mechanics. *Z Wahrscheinlichkeitstheorie Verw Gebiete*, 1978, 44: 117–139
- 10 Ellis R S, Newman C M. Fluctuations in Curie-Weiss exemplis. In: Dell'Antonio G, Doplicher S, Jona-Lasinio G, eds. *Mathematical Problems in Theoretical Physics*. Lecture Notes in Physics, vol. 80. Berlin-Heidelberg: Springer, 1978, 313–324
- 11 Ellis R S, Newman C M. The statistics of Curie-Weiss models. *J Stat Phys*, 1978, 19: 149–161
- 12 Ellis R S, Newman C M, Rosen J S. Limit theorems for sums of dependent random variables occurring in statistical mechanics. *Probab Theory Related Fields*, 1978, 44: 153–169
- 13 Goldstein L, Reinert G. Stein's method for the beta distribution and the pólya-eggenberger urn. *J Appl Probab*, 2013, 50: 1187–1205
- 14 Kusuoka S, Tudor C A. Stein method for invariant measures of diffusions via Malliavin calculus. *Stochastic Process Appl*, 2012, 122: 1627–1651
- 15 Luk H M. Stein's method for the Gamma distribution and related statistical applications. PhD Thesis. Ann Arbor: University of Southern California, 1994
- 16 Meckes M W, Meckes E S. The central limit problem for random vectors with symmetries. *J Theoret Probab*, 2007, 20: 697–720
- 17 Papangelou F. On the Gaussian fluctuations of the critical curie-weiss model in statistical mechanics. *Probab Theory Related Fields*, 1989, 83: 265–278
- 18 Stein C. A bound for the error in the normal approximation to the distribution of a sum of dependent random variables. In: *Proceedings of the Sixth Berkeley Symposium on Mathematical Statistics and Probability*. Probability Theory, vol. 2. Berkeley: University of California Press, 1972, 583–602